

Kernelized Covariance for Action Recognition

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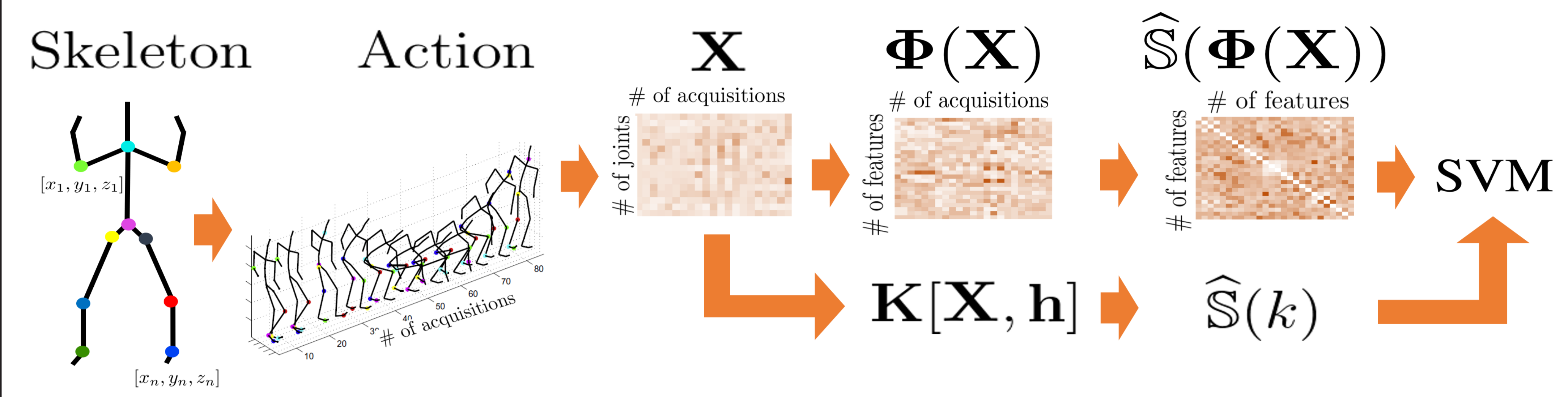
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Abstract

- Originally devised as an image descriptor [5], the covariance matrix is powerful in correlating skeletal joints across time for action recognition [2, 10].
- As the main limitation, covariance can only capture *linear* mutual relationships.
- In this work, we extend **covariance to model arbitrary, non-linear relationships** by recovering the applicability of the **kernel trick** and consequently **avoiding any preliminary feature encoding** of the raw data.

Pipeline



Recovering the Kernel Trick for the Covariance Representation

Each action instance is represented through $\mathbf{X} = [\mathbf{x}(1), \dots, \mathbf{x}(T)] \in \mathbb{R}^{3n \times T}$, collecting the n joints positions $\mathbf{x}_1(t), \mathbf{x}_2(t), \dots, \mathbf{x}_n(t)$ at each timestamp $t = 1, \dots, T$, being $\mathbf{x}_i(t) = [x_i(t) y_i(t) z_i(t)]$. We introduce the sampling covariance operator, defined as

$$\hat{\mathbf{S}}(\mathbf{X}) = \frac{1}{T-1} \sum_{t=1}^T (\mathbf{x}(t) - \boldsymbol{\mu})(\mathbf{x}(t) - \boldsymbol{\mu})^\top, \text{ where } \boldsymbol{\mu} = \frac{1}{T} \sum_{s=1}^T \mathbf{x}(s), \text{ which rewrites } \hat{\mathbf{S}}(\mathbf{X}) = \mathbf{X}\mathbf{P}\mathbf{X}^\top \text{ once defined } \mathbf{P} = \frac{1}{T-1} \left(\frac{1}{T} \mathbf{I} - \mathbf{1} \right)$$

As to model non-linear correlations within the data through covariance, $\hat{\mathbf{S}}(\Phi(\mathbf{X})) = \Phi(\mathbf{X})\mathbf{P}\Phi(\mathbf{X})^\top$ is computed in terms of an explicit feature map $\Phi: \mathbb{R}^{3n} \rightarrow \mathcal{H}$ which has to be applied to the whole data matrix \mathbf{X} , obtaining $\Phi(\mathbf{X}) = [\Phi(\mathbf{x}(1)), \dots, \Phi(\mathbf{x}(T))] \in \mathcal{H} \times \dots \times \mathcal{H}$.

We posit that $\hat{\mathbf{S}}(\Phi(\mathbf{X}))$ can be computed in terms of the kernel $k(\mathbf{x}, \mathbf{y}) = \langle \Phi(\mathbf{x}), \Phi(\mathbf{y}) \rangle_{\mathcal{H}}$ **only**, if provided the existence of \mathbf{h}_i such that $\Phi(\mathbf{h}_i) = \mathbf{e}_i$ for any i . Indeed,

$$\hat{\mathbf{S}}_{ij}(\Phi(\mathbf{X})) = \sum_{s=1}^T \sum_{t=1}^T \Phi_i(\mathbf{x}(s)) \mathbf{P}_{st} \Phi_j(\mathbf{x}(t)) = \sum_{s=1}^T \sum_{t=1}^T \langle \Phi(\mathbf{x}(s)), \mathbf{e}_i \rangle_{\mathcal{H}} \mathbf{P}_{st} \langle \Phi(\mathbf{x}(t)), \mathbf{e}_j \rangle_{\mathcal{H}} = \sum_{s=1}^T \sum_{t=1}^T \underbrace{\langle \Phi(\mathbf{x}(s)), \Phi(\mathbf{h}_i) \rangle_{\mathcal{H}}}_{k(\mathbf{x}(s), \mathbf{h}_i)} \mathbf{P}_{st} \underbrace{\langle \Phi(\mathbf{x}(t)), \Phi(\mathbf{h}_j) \rangle_{\mathcal{H}}}_{k(\mathbf{x}(t), \mathbf{h}_j)} = \hat{\mathbf{S}}_{ij}(k).$$

As the main theoretical contribution of our work, we show that the assumption $\Phi(\mathbf{h}_i) = \mathbf{e}_i$ can be fulfilled by a particular class of random feature maps related to a Taylor kernel function $k(\mathbf{x}, \mathbf{z}) = \sum_{\ell=0}^{\infty} a_{\ell} \langle \mathbf{x}, \mathbf{z} \rangle^{\ell}$, being $a_{\ell} > 0$ for any ℓ . For concreteness, let us fix $k(\mathbf{x}, \mathbf{z}) = \exp(\gamma \cdot \langle \mathbf{x}, \mathbf{z} \rangle)$ for a given $\gamma > 0$.

Random Approximated Feature Map [3]

$\Psi: \mathbb{R}^{3n} \rightarrow \mathcal{H}$, $\Psi_i(\mathbf{x}) \stackrel{i.i.d.}{\sim} \sqrt{a_N p^{N+1}} \prod_{j=1}^N \langle \omega_j, \mathbf{x} \rangle$, where N is sampled with prob. $1/p^{N+1}$ for the parameter $p > 1$ and $\omega_1, \dots, \omega_N$ are Rademacher distributed. Then,

$$k(\mathbf{x}, \mathbf{z}) = \sum_{i=0}^{\infty} \Psi_i(\mathbf{x}) \Psi_i(\mathbf{z}) \approx \sum_{i=0}^M \Psi_i(\mathbf{x}) \Psi_i(\mathbf{z})$$

where the approximation holds in mean over ω_j and uniformly in concentration for \mathbf{x}, \mathbf{z} lying in a compact set. \Rightarrow for the experiments $M = 3n$

Kernelizing the Covariance Representation

For $i = 1, \dots, M$ we can **compute**^a \mathbf{h}_i such that

$$\Psi|_M(\mathbf{h}_i) = \frac{1}{\sqrt{M}} [\Psi_1(\mathbf{h}_i), \dots, \Psi_M(\mathbf{h}_i)] = \mathbf{e}_i.$$

Therefore, once defined $K_{is}[\mathbf{X}, \mathbf{h}] = k(\mathbf{x}(s), \mathbf{h}_i)$,

$$\hat{\mathbf{S}}(k) = \mathbf{K}[\mathbf{X}, \mathbf{h}] \mathbf{P} \mathbf{K}[\mathbf{X}, \mathbf{h}]^\top = \hat{\mathbf{S}}(\Psi|_M(\mathbf{X})) \approx \hat{\mathbf{S}}(\Psi(\mathbf{X})) = \hat{\mathbf{S}}(\Phi(\mathbf{X})).$$

^asee Proposition 1. in the paper

The Pseudocode

1. For each action, extract the data matrix \mathbf{X} collecting all the T temporal observations $\mathbf{x}(1), \dots, \mathbf{x}(T)$, each one encoding the 3D coordinates of the n joints.
2. For each data matrix \mathbf{X} , select $\mathbf{h}_1, \dots, \mathbf{h}_M$ as in Proposition 1. and compute $\mathbf{K}[\mathbf{X}, \mathbf{h}]$.
3. Compute the linear operator \mathbf{P} .
4. By means of $\mathbf{K}[\mathbf{X}, \mathbf{h}]$ and \mathbf{P} , compute $\hat{\mathbf{S}}(k)$.

Experimental Results

Publicly available code at <https://www.iit.it/pavis/code/kcar>

For any action trial \mathbf{X} , compute the kernel matrix $\mathbf{K}[\mathbf{X}, \mathbf{h}]$ through $k(\mathbf{x}(s), \mathbf{h}_i) = \exp(\gamma \cdot \langle \mathbf{x}(s), \mathbf{h}_i \rangle)$ for $s = 1, \dots, T$ and $i = 1, \dots, 3n$. The final classification step is performed with a support vector machine (SVM) fed with a log-Euclidean Gaussian kernel. The SVM cost value C and the kernel parameter γ are cross-validated.

Method	MSR-Action3D
Action Graph [4]	79.0%
Random Occupancy Patterns [8]	86.5%
Actionlets [9]	88.2%
Pose Set [7]	90.0%
Moving Pose [12]	91.7%
Lie Group [6]	92.5%
Normal Vectors [11]	93.1%
Kernelized-COV (proposed)	96.2%

Method	MSR-Action3D	MSR-Daily-Activity	MSRC-Kinect12	HDM-05
Region-COV [5]	74.0%	85.0%	89.2%	91.5%
Hierarchy of COVs [2]	90.5%	—	91.7%	—
COV- $J_{\mathcal{H}}$ -SVM [1]	80.4%	75.5%	89.2%	82.5%
Ker-RP-POL [10]	96.2%	96.9%	90.5%	93.6%
Ker-RP-RBF [10]	96.9%	96.3%	92.3%	96.8%
Kernelized-COV (proposed)	96.2%	96.3%	95.0%	98.1%

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